Generating random variables II (Exercises)

Anatoli Iambartsev IME-USP **Exercise 1.**^{*} Simulate random bi-dimensional vector (ξ, η) with density p(x, y) = 6x on the triangle

$$T = \{(x, y) : x > 0, y > 0, x + y \le 1\}.$$

*[S], p.56, chapter 2

The decomposition

$$p_{\xi\eta}(x,y) = p_{\xi}(x)p_{\eta|\xi}(y \mid x) = p_{\eta}(x)p_{\xi|\eta}(x \mid y),$$

provides two cases: (a) generate first ξ second η ; (b) first η second ξ .

(a) Consider the first case:

$$p_{\xi}(x) = \int_{0}^{1-x} p_{\xi\eta}(x, y) dy = 6x(1-x), \ 0 < x < 1,$$

$$p_{\eta|\xi}(y \mid x) = p_{\xi\eta}(x, y) / p_{\xi}(x) = (1-x)^{-1}, \ 0 < y < 1-x$$

$$p_{\xi}(x) = 6x(1-x), x \in [0,1], p_{\eta|\xi}(y \mid x) = (1-x)^{-1}, y \in [0,1-x].$$

$$F_{\xi}(x) = \int_{0}^{x} p_{\xi}(u) du = 3x^{2} - 2x^{3}, \ x \in [0, 1],$$

$$F_{\eta|\xi}(y \mid x) = \int_{0}^{y} p_{\eta|\xi}(u \mid x) du = y(1 - x)^{-1}, \ y \in [0, 1 - x],$$

Thus, let $U_1, U_2 \sim U[0, 1]$ then ξ, η is the solution of

$$\begin{cases} 3\xi^2 + 2\xi^3 = U_1 \\ \eta = U_2(1 - \xi) \end{cases}$$

(b) first η , second ξ

$$p_{\eta}(y) = \int_{0}^{1-y} p_{\xi\eta}(x,y) dx = 3(1-y)^{2}, \ y \in (0,1)$$
$$p_{\xi|\eta}(x \mid y) = p_{\xi\eta}(x,y)/p_{\eta}(y) = 2x(1-y)^{-2}, \ x \in (0,1-y)$$

$$F_{\eta}(y) = \int_{0}^{y} p_{\eta}(v) dv = 1 - (1 - y)^{3}, \ y \in (0, 1)$$

$$F_{\xi|\eta}(x \mid y) = \int_{0}^{x} p_{\xi|\eta}(v \mid y) dv = x^{2}(1 - y)^{-2}, \ x \in (0, 1 - y)$$

$$F_{\eta}(y) = 1 - (1 - y)^3, \ y \in (0, 1)$$

 $F_{\xi|\eta}(x \mid y) = x^2(1 - y)^{-2}, \ x \in (0, 1 - y)$
Thus

$$\begin{cases} U_1 = 1 - (1 - \eta)^3 \\ U_2 = \xi^2 (1 - \eta)^{-2} \end{cases}$$

$$\begin{cases} U_1 = 1 - (1 - \eta)^3 \\ U_2 = \xi^2 (1 - \eta)^{-2} \end{cases} \Longrightarrow \begin{cases} U_1 = (1 - \eta)^3 \\ U_2 = \xi^2 (1 - \eta)^{-2} \end{cases}$$

Finally

$$\implies \begin{cases} \eta = 1 - (U_1)^{1/3} \\ \xi = (U_2)^{1/2} (U_1)^{1/3} \end{cases}$$

the second variant is better, because we need to solve the equation of 3rd order in the first variant.

$$B = \{x^2 + y^2 + z^2 < R^2\}.$$

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Let $Q = (\xi, \eta, \zeta)$. The joint density is constant in the ball

$$p_Q(x, y, z) = \frac{1}{(4/3)\pi R^3}, \ (x, y, z) \in B.$$

Direct calculation of conditional densities is boring. Let us try to change coordinates.

$$B = \{x^2 + y^2 + z^2 < R^2\}.$$

In polar coordinates

$$x = r \sin \theta \cos \phi, y = r \sin \theta \sin \phi, z = r \cos \theta.$$

Jacobian

$$\left|\frac{\partial(x,y,z)}{\partial(r,\theta,\phi)}\right| = r^2 \sin\theta$$

and

$$p_Q(r,\theta,\phi) = \frac{1}{(4/3)\pi R^3} r^2 \sin\theta$$

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$$p_Q(r,\theta,\phi) = \frac{1}{(4/3)\pi R^3} r^2 \sin\theta$$
$$= \frac{3r^2}{R^3} \times \frac{\sin\theta}{2} \times \frac{1}{2\pi}$$

It means that coordinates r_Q, θ_Q, ϕ_Q of point Q are independents.

$$B = \{x^2 + y^2 + z^2 < R^2\}.$$

Equations for inverse cdf are

$$\int_0^{r_Q} \frac{3r^2}{R^3} dr = U_1, \int_0^{\theta_Q} \frac{\sin\theta}{2} d\theta = 1 - U_2, \int_0^{\phi_Q} \frac{1}{2\pi} d\phi.$$

Thus we obtain

$$r_Q = R(U_1)^{1/3}, \cos \theta_Q = 2U_2 - 1, \phi_Q = 2\pi U_3.$$

Integral superposition method.

Generalization of the superposition method where the target distribution F(x) can be represented as

$$F(x) = \sum_{k} c_k F_k(x).$$

Integral superposition method based on

$$p_{\xi}(x) = \int_{-\infty}^{\infty} p(x, y) dy.$$

It is rarely used. It is useful when a density represented as an integral.

Integral superposition method.

Example. Let

$$p_{\xi}(x) = n \int_{1}^{\infty} y^{-n} e^{-xy} dy$$

use $p(x,y) = p_{\eta}(y)p_{\xi}(x \mid y)$.

$$p_{\eta}(y) = \int_{0}^{\infty} p(x, y) dx = n y^{-n-1}, \ 1 < y < \infty,$$

 $p_{\xi}(x \mid y) = p(x, y) / p_{\eta}(y) = y e^{-xy}, \ 0 < x < \infty,$

$$\implies F_{\eta}(y) = 1 - y^{-n}, \ F_{\xi}(x \mid y) = 1 - e^{-xy}$$
$$\implies \eta = (U_1)^{-1/n}, \ \xi = -\frac{\ln U_2}{\eta}$$
$$\implies \xi = -(U_1)^{1/n} \ln U_2.$$

Exercise 3. Let $X \sim F(x) = x^n, x \in [0, 1]$. By inverse method $X = (U)^{1/n}$. Prove that X can be simulated by $X = \max\{U_1, \ldots, U_n\}$.

Exercise 4. Let $X \sim \Gamma(n)$: the density $p_n(x) = (\Gamma(n))^{-1}x^{n-1}e^{-x}$, where $n \ge 1$ is integer. Prove that for any n r.v. X can be generated as $X = -\ln(U_1 \cdots U_n)$.

Exercise 5. Let $X \sim B(n,p)$. Prove $X = \sum_{i=1}^{n} [p - U_i]_+$, where $[x]_+ = 1$ if $x \ge 0$ and $[x]_+ = 0$ if x < 0.

Exercise 6. Simulate $X \sim F(x) = 1 - \frac{1}{3}(2e^{-x} + e^{-5x}), x \in (0, \infty)$ using inverse method.

Exercise 7. Simulate point P = (X, Y) uniformly distributed in anel $R_1^2 < x^2 + y^2 < R_2^2$.

Exercise 8*. Energy of a neutron is considered as a r.v. $X \in (0, \infty)$ with density

$$p(x) = c e^{-x/T} \mathrm{sh}\big(b\sqrt{2x/T}\big),$$

where b,T are parameters, and c is normalising constant $c = \sqrt{2/\pi}e^{-b^2/2}(bT)^{-1}$. Prove that ξ can be represented as

$$\xi = T\left(\frac{(\zeta+b)^2}{2} - \ln U\right),\,$$

where $\zeta \sim N(0, 1)$ and ζ, U are independent.

Exercise 9. Suggest the simulation for the r.v.

$$X \sim F(t) = \exp\left\{-\frac{cF}{\kappa}t\right\} \exp\left\{\frac{cV}{(\gamma\kappa)^2n}(1-e^{-\gamma\kappa nt})\right\}$$
$$\times c\left(\frac{F}{\kappa}-\frac{V}{\gamma\kappa}e^{-\gamma\kappa nt}\right).$$

where c, F, V, κ, γ are parameters.

References:

- [RC] Cristian P. Robert and George Casella. Introducing Monte Carlo Methods with R. Series "Use R!". Springer
 - [S] Sobol, I.M. *Monte-Carlo numerical methods.* Nauka, Moscow, 1973. (In Russian)