

# The FBST for Model Selection in Mixture of Multivariate Normals

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The FBST Value of Evidence  
Full Bayesian Significance Test  
(Pereira and Stern, 1999)

Posterior density, likelihood and prior:

$$p_x(\theta) \propto L(\theta | x) p(\theta).$$

Null hypothesis:

$$\Theta_H = \{\theta \in \Theta \mid g(\theta) \leq \mathbf{0} \wedge h(\theta) = \mathbf{0}\}$$

Sharp (precise) hypotheses:  
 $\dim(\Theta_H) < \dim(\Theta)$ .

Evidence against the hypothesis:

$$\begin{aligned}\text{Ev}(H) &= \int_{T_H} p_x(\theta) d\theta , \text{ where} \\ T_H &= \{\theta \in \Theta \mid s(\theta) > s_H\} \\ s_H &= \sup_{\theta \in \Theta_H} s(\theta) \\ s(\theta) &= \left( \frac{p_x(\theta)}{r(\theta)} \right)\end{aligned}$$

$s(\theta)$  is the Posterior Surprise

If the reference density  $r(\theta) \propto 1$  ,  
the Tangent set  $T_H$  , HRSS = HDPS

Operationally:

Optimization + Integration step

## Mixture Models

Sample:  $x^j$  ,  $j = 1 \dots n$

Coord:  $x_h^j$  ,  $h = 1 \dots d$

Classes:  $c(j) = k$  ,  $k = 1 \dots m$

Latent Variables:  $z_k^j = \mathbf{1} (c(j) = k)$

Numb. samp. class  $k$ ,  $y_k$  :  $y = Z\mathbf{1}$

Mixture parameters:

$$\Pr(c(j) = k) = w_k$$

$$\text{if } c(j) = k \text{ then } x^j \sim f(x^j | \psi_k)$$

$$\theta = [w_1, \dots, w_k, \psi_1, \dots, \psi_k]$$

Conditional on the missing data:

$$\begin{aligned} f(x^j | \theta) &= \sum_{k=1}^m f(x^j | \theta, z_k^j) f(z_k^j | \theta) \\ &= \sum_{k=1}^m w_k f(x^j | \psi_k) \\ f(X | \theta) &= \prod_{j=1}^n f(x^j | \theta) \\ &= \prod_{j=1}^n \sum_{k=1}^m w_k f(x^j | \psi_k) \end{aligned}$$

Conditional classification probabilities,

$P = f(Z | X, \theta)$ :

$$\begin{aligned} p_k^j &= f(z_k^j | x^j, \theta) = \frac{f(z_k^j, x^j | \theta)}{f(x^j | \theta)} \\ &= \frac{w_k f(x^j | \psi_k)}{\sum_{k=1}^m w_k f(x^j | \psi_k)} \end{aligned}$$

Likelihood for the “completed” data,  $X, Z$ :

$$\begin{aligned} f(X, Z | \theta) &= \prod_{j=1}^n f(x^j | \psi_{c(j)}) f(z_k^j | \theta) \\ &= \prod_{k=1}^m \left[ (w_k)^{y_k} \prod_{j | c(j)=k} f(x^j | \psi_k) \right] \end{aligned}$$

where  $y_k = \sum_j z_k^j$ .

## Normal-Wishart Distribution

$u$  and  $S$  are the statistics:

$$\begin{aligned} u &= \frac{1}{n} \sum_{j=1}^n x^j = \frac{1}{n} X \mathbf{1} \\ S &= \sum_{j=1}^n (x^j - b) \otimes (x^j - b)' \\ &= (X - b)(X - b)' \end{aligned}$$

$u$  Normal, mean  $b$ , precision  $nR$ .

$S$  Wishart,  $n$  d.o.freedom, precision  $R$ .

$$\begin{aligned} N(u | n, b, R) &= \left(\frac{n}{2\pi}\right)^{d/2} |R|^{1/2} \\ &\quad \exp\left(-\frac{n}{2}(u - b)'R(u - b)\right) \\ W(S | e, R) &= c^{-1} |S|^{(e-d-1)/2} \\ &\quad \exp\left(-\frac{1}{2}\text{tr}(S R)\right) \end{aligned}$$

$X$ , unknown mean and precision,  $b$ ,  $R$

$$\begin{aligned} u &= (1/n)X\mathbf{1} \\ S &= (X - u)(X - u)' \end{aligned}$$

Posterior Normar-Wishart distribution:

$$\begin{aligned} NW(b, R | \ddot{n}, \ddot{e}, \ddot{u}, \ddot{S}) \\ &= W(R | \ddot{e}, \ddot{S}) N(b | \ddot{n}, \ddot{u}, R) \\ \ddot{n} &= \dot{n} + n \\ \ddot{e} &= \dot{e} + n \\ \ddot{u} &= (nu + \dot{n}\dot{u})/\ddot{n} \\ \ddot{S} &= S + \dot{S} + \frac{n\dot{n}}{n + \dot{n}}(u - \dot{u}) \otimes (u - \dot{u})' \end{aligned}$$

One dot  $\Rightarrow$  Prior parameters

Two dots  $\Rightarrow$  Posterior parameters

Non-informative parameters:

$$\dot{n} = 0, \dot{u} = 0, \dot{e} = 0, \dot{S} = 0.$$

Dirichlet-Multinomial distribution:

$$M(y | n, w) = \frac{n!}{y_1! \dots y_m!} (w_1)^{y_1} \dots (w_m)^{y_m}$$

$$D(w | y) = \frac{\Gamma(y_1 + \dots + y_k)}{\Gamma(y_1) \dots \Gamma(y_k)} \prod_{k=1}^m w_k^{y_k - 1}$$

$w > \mathbf{0}$ ,  $w\mathbf{1} = \mathbf{1}$ .

Posterior:  $\ddot{y} = \dot{y} + y$ .

Non-informative prior:  $\dot{y} = \mathbf{1}$ .

Finally, Dirichlet-Normal-Wishart posterior,

$$f(\theta | X, \dot{\theta}) = f(X | \theta) f(\theta | \dot{\theta})$$

$$f(X | \theta) = \prod_{j=1}^n \sum_{k=1}^m p_k^j w_k N(x^j | b^k, R^k)$$

$$f(\theta | \dot{\theta}) = D(w | \dot{y}) \prod_{k=1}^m NW(b^k, R^k | \dot{n}_k, \dot{e}_k, \dot{u}^k, \dot{S}^k)$$

$$p_k^j = \frac{w_k N(x^j | b^k, R^k)}{\sum_{k=1}^m w_k N(x^j | b^k, R^k)}$$

and completed posterior:

$$\begin{aligned}
f(\theta | X, Z, \dot{\theta}) &= f(\theta | X, Z) f(\theta | \dot{\theta}) = \\
&= D(w | \ddot{y}) \prod_{k=1}^m NW(b^k, R^k | \ddot{n}_k, \ddot{e}_k, \ddot{u}^k, \ddot{S}^k) \\
y &= Z\mathbf{1} , \quad \ddot{y} = \dot{y} + y \\
\ddot{n} &= \dot{n} + y , \quad \ddot{e} = \dot{e} + y \\
u^k &= \frac{1}{y_k} \sum_{j=1}^n z_k^j x^j \\
S^k &= \sum_{j=1}^n z_k^j (x^j - u^k) \otimes (x^j - u^k)' \\
\ddot{u}^k &= \frac{\dot{n}_k \dot{u}^k + y_k u^k}{\ddot{y}_k} \\
\ddot{S}^k &= S^k + \dot{S}^k + \frac{\dot{n}_k y_k}{\ddot{n}_k} (u^k - \dot{u}^k) \otimes (u^k - \dot{u}^k)'
\end{aligned}$$

Model: D-M-N-W mixture,  $d = 2, m = 2$ ,

FBST for Model Selection:  $H : m = 1$

Optimization step:

Local: EM or Box-Quacan

Global: MCMC + Cluster Filter, SEM, etc.

Integration step: MCMC

$$\begin{aligned} f(z^j | p^j) &= M(z^j | 1, p^j) \\ f(w | Z, \dot{y}) &= D(w | \dot{y}) \\ f(R^k | X, Z, \dot{e}_k, \dot{S}^k) &= W(R | \ddot{e}_k, \ddot{S}^k) \\ f(b^k | X, Z, R^k, \dot{n}_k, \dot{u}^k) &= N(b | \ddot{n}_k, \ddot{u}^k, R^k) \end{aligned}$$

Label Switching:  $\text{perm}([1 \dots m])$

Break all non-identifiability symmetries.

Ex: Order components by linear combination of vector means,  $c' b^k$ .

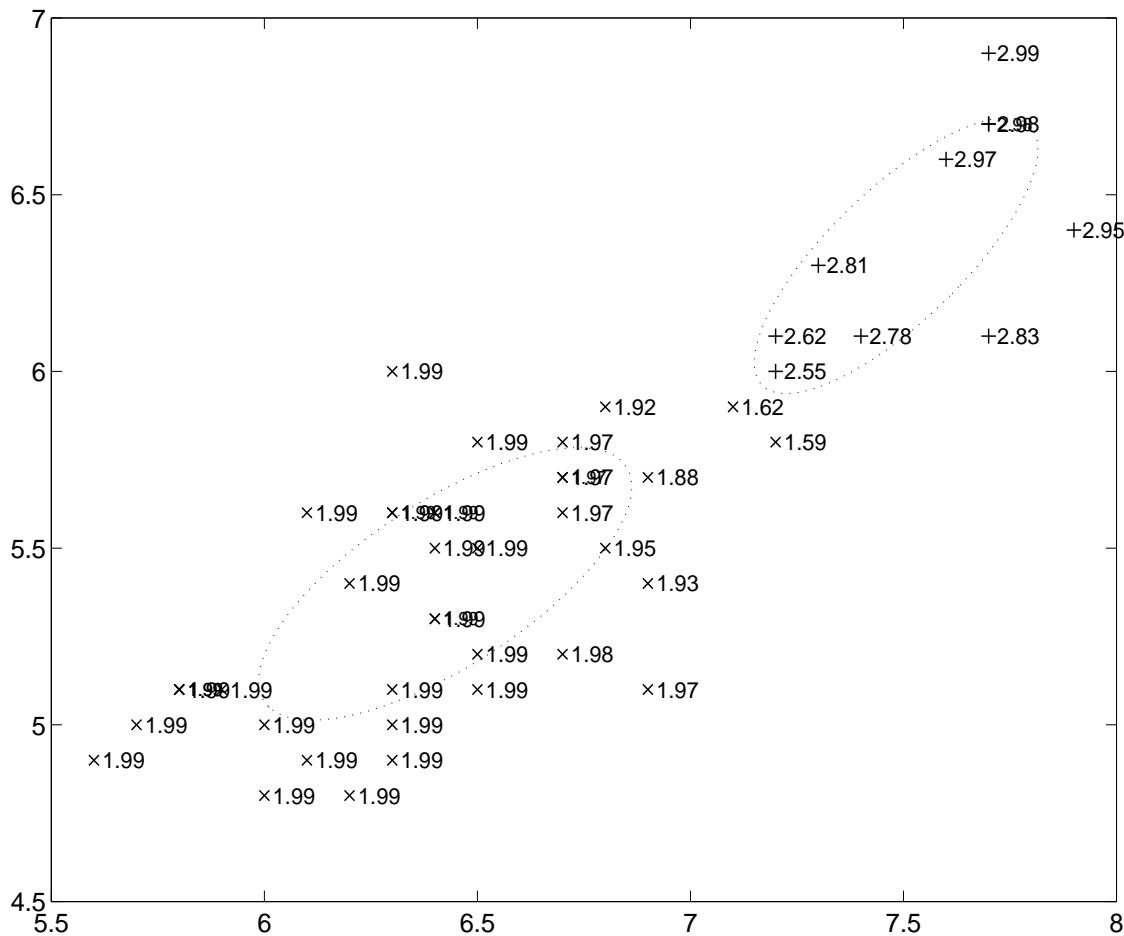
Avoid Trapping states, near-singular  $V^k$

- Vague Priors: Empirical approx.

Forbidden States: - Rejection rules

How many components in the Iris data?

Sepal + pedal length, 49 points

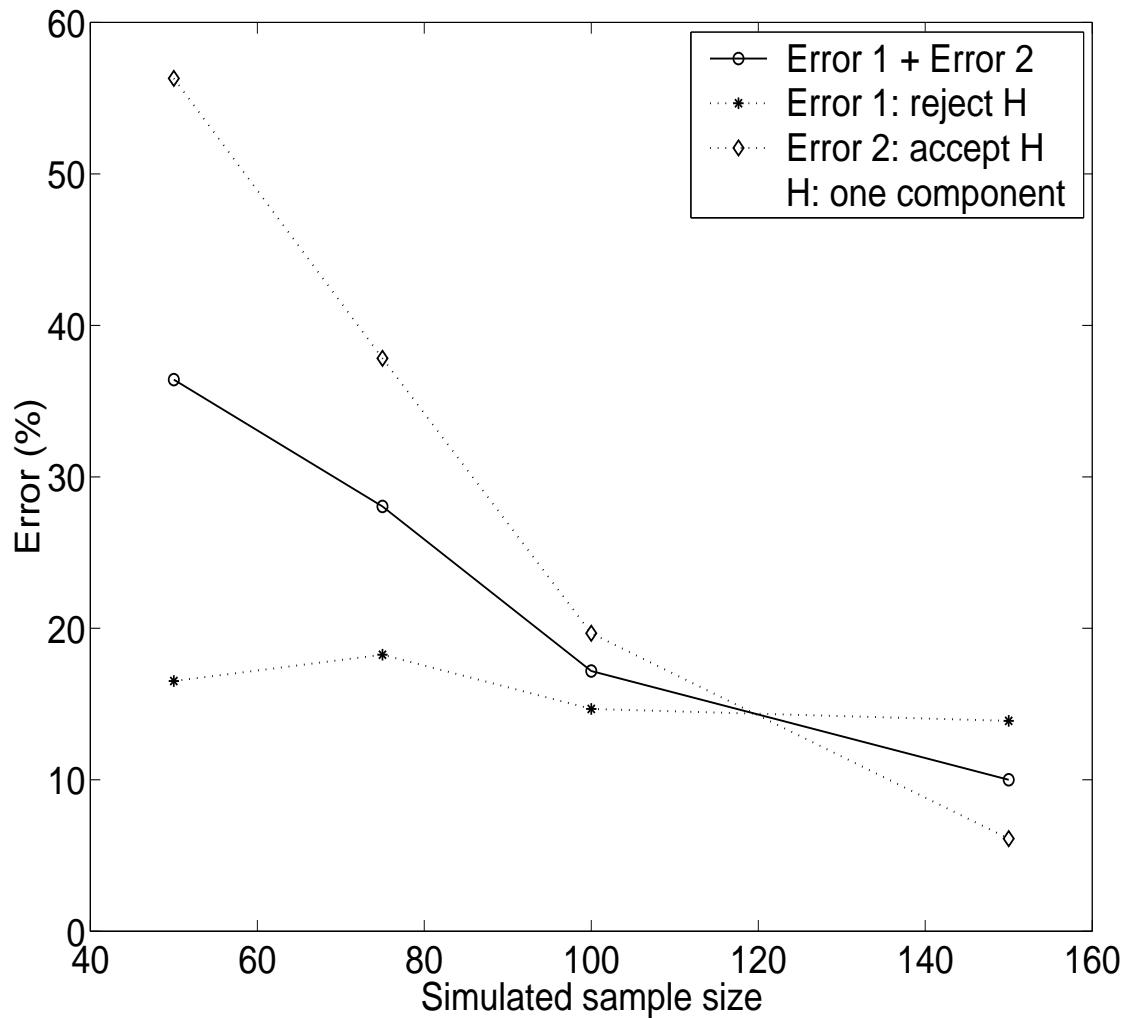


Analysis of empirical errors for FBST:

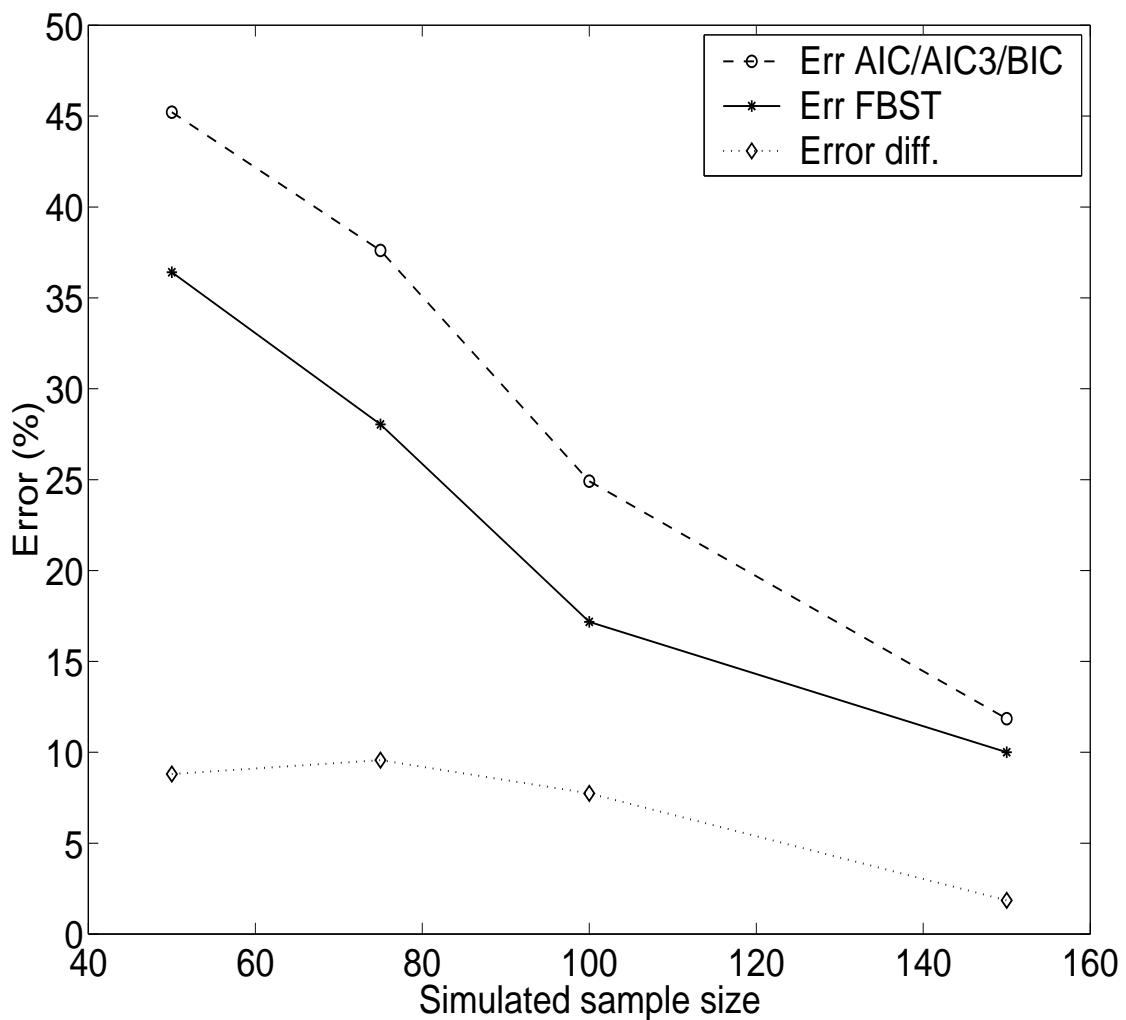
Sample sizes  $n = 50, 75, 100, 150$

$\alpha$ : type 1 error ,  $\beta$ : type 2 error

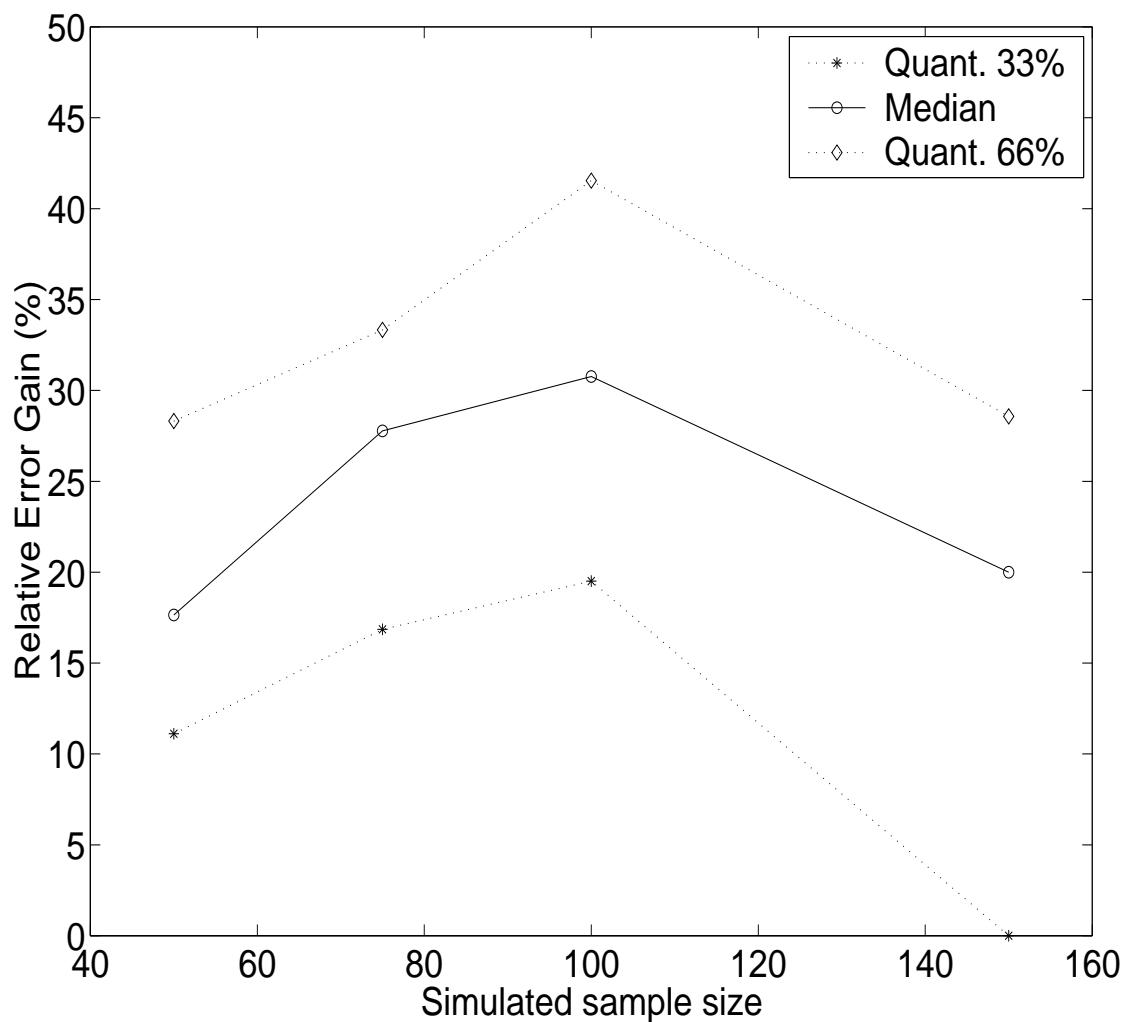
Total error:  $\alpha + \beta$



Comparison of empirical errors:  
Better of AIC, AIC3, BIC vs. FBST  
Sample sizes  $n = 50, 75, 100, 150$



## Media, 0.33 and 0.66 quantiles for error rates differences



Conclusions:

FBST: Smaller error in model selection  
(up to 35%) for small ( $< 150$ ) samples  
but computationally more expensive

AIC: Same error for large ( $> 150$ ) samples  
and computationally cheaper

Further Research:

Mixture of components of different type,

$$c(j) = k \Rightarrow x^j \sim f_k(x^j | \psi_k)$$

Tests for Separate Hypotheses.

## Appendix: Critical evidence for rejecting $H$

Given a sample  $X_0$ , we must establish a critical level  $cl$  such that

if  $Ev(H) > cl$  then reject  $H$

Optimal parameters:

$$\begin{aligned}\theta^* &= \arg \max_{\theta \in \Theta_H} f(\theta | X_0) \\ &= [w^*, b^{k*}, R^{k*}] \\ \hat{\theta} &= \arg \max_{\theta \in \Theta} f(\theta | X_0) \\ &= [\hat{w}, \hat{b}^k, \hat{R}^k]\end{aligned}$$

Simulation of new samples  $\{{}_l^1 X\}$  and  $\{{}_l^2 X\}$ :

$$\begin{aligned} f({}_l^1 z^j | \theta^*) &= M({}_l^1 z^j | 1, w^*) \\ f({}_l^1 x^j | {}_l^1 z^j, \theta^*) &= N({}_l^1 x^j | b^{k*}, R^{k*}) {}_l^1 z^j \\ f({}_l^2 z^j | \hat{\theta}) &= M({}_l^2 z^j | 1, \hat{w}) \\ f({}_l^2 x^j | {}_l^2 z^j, \hat{\theta}) &= N({}_l^2 x^j | \hat{b}^k, \hat{R}^k) {}_l^2 z^j \end{aligned}$$

Type 1 error ( $\alpha$ ) is computed over  $\{{}_l^1 X\}$

Type 2 error ( $\beta$ ) is computed over  $\{{}_l^2 X\}$

Calibrate  $cl$  that minimizes  $\alpha + \beta$ .

Small run  $l = 1, \dots, r$ , poor calibration.