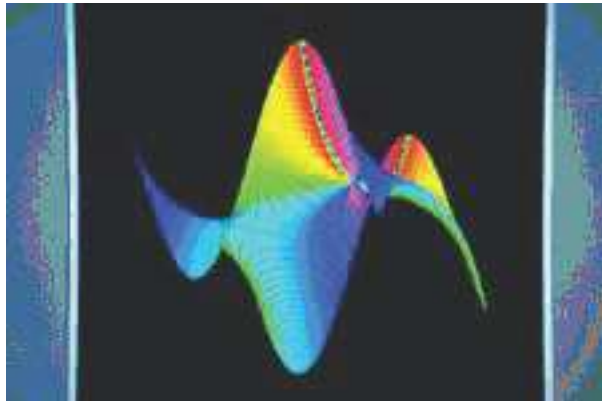


Constrained Pattern Matching*

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Outline

1. Pattern Matching and Constrained Pattern Matching Problems
2. Combinatorial Approach and Language Representation
3. Number of Pattern Occurrences and Analytical Results
4. Experimental Results
5. Proof Sketch of Large Deviation Result

Pattern Matching

Let \mathcal{W} and T be (set of) strings generated over a finite alphabet \mathcal{A} .

We call \mathcal{W} the **pattern** and T the **text**. The text T is of length n and is generated by a **probabilistic source**.

The pattern \mathcal{W} can be a single string

$$\mathcal{W} = w_1 \dots w_m, \quad w_i \in \mathcal{A}$$

or a set of strings

$$\mathcal{W} = \{\mathcal{W}_1, \dots, \mathcal{W}_d\}$$

with $\mathcal{W}_i \in \mathcal{A}^{m_i}$ being a set of strings of length m_i .

Questions

- How many times does \mathcal{W} occur in T ?
- What is the **probability** that \mathcal{W} occurs exactly r times in T ?

Constrained Pattern Matching

There are **constraints** on the text T . (e.g., (d, k) sequences, regular expression)

A (d, k) sequence is a **binary sequence** in which any **run of zeros** must be of length **at least d** and **at most k** .

Example: $(2, 4)$ sequence - 0010001001000100001000100100001000

(d, k) sequences are useful for **digital recording** and **biology**.



Questions

- How many times does W occur in a (d, k) sequence, T ?
- What is the **conditional probability** that W occurs exactly r times in a (d, k) sequence, T ?

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Combinatorial Approach

We use a **combinatorial approach**, based on (M. Régnier & W. Szpankowski, *Algorithmica*, 1998), (P. Jacquet & W. Szpankowski, *ISIT*, 2006).

- Construct **languages** and their **relationships**
- Translate the language relationships into **generating functions**

A language, say \mathcal{L} , is a **collection of words**, and its **probability generating function** is defined as

$$L(z) = \sum_{u \in \mathcal{L}} P(u) z^{|u|} = \sum_{n \geq 0} z^n L_n, \quad [z^n] L(z) = L_n$$

where $P(u)$ is the **probability of u** .

Define

$$\mathcal{A}_{d,k} = \{\underbrace{0 \dots 0}_d, \dots, \underbrace{0 \dots 0}_k\},$$

that is, a set of **runs of zeros of length between d and k** .

Combinatorial Approach

Define

$$\mathcal{B}_{d,k} = \mathcal{A}_{d,k} \cdot \{1\} = \{\underbrace{0 \dots 0}_d 1, \dots, \underbrace{0 \dots 0}_k 1\}$$

as an *extended alphabet*.

The probability generating function of $\mathcal{B}_{d,k}$ is

$$B(z) = p^d q z^{d+1} + p^{d+1} q z^{d+2} + \dots + p^k q z^{k+1} = zq \frac{(zp)^d - (zp)^{k+1}}{1 - zp},$$

where p is the probability of emitting a '0' and $q = 1 - p$.

We consider only *restricted* (d, k) sequences, which are (d, k) sequences that start with '0' and end with '1'.

Observe that the set of all *restricted* (d, k) sequences is

$$\mathcal{B}_{d,k}^* = \{\epsilon\} + \mathcal{B}_{d,k} + \mathcal{B}_{d,k}^2 + \mathcal{B}_{d,k}^3 + \dots, \text{ and } B^*(z) = \frac{1}{1 - B(z)}.$$

Note: We only consider occurrences of the pattern w over $\mathcal{B}_{d,k}$, not over the binary alphabet.

Example: $w = 01$ occurs only once in a sequence 001010001.

Autocorrelation Set

Let $w = \beta_1 \dots \beta_m$, where $\beta_i \in \mathcal{B}_{d,k}$.

We define the *autocorrelation set* of w over $\mathcal{B}_{d,k}$ as

$$\mathcal{S} = \{\beta_{l+1}^m : \beta_1^l = \beta_{m-l+1}^m\}, \quad 1 \leq l \leq m$$

where $\beta_i^j = \beta_i \dots \beta_j$. Its probability generating function $S(z)$ is called the *autocorrelation polynomial*. (as in L. Guibas & A.M. Odlyzko, 1981)

Example: Let $w = 0100101$ over $\mathcal{B} = \{01, 001, 0001\}$.

Then

$$\mathcal{S} = \{\varepsilon, 00101\}$$

since

$$\begin{array}{ccc} 01 & 001 & 01 \\ & & 01 \quad 001 \quad 01. \end{array}$$

Note that $S(z) = 1 + P(00101)z^5$.

Language \mathcal{T}_r

\mathcal{T}_r – the set of all **restricted** (d, k) **sequences** containing exactly r occurrences of w . (M. Régnier and W. Szpankowski, 1998)

We define some languages: \mathcal{R} , \mathcal{U} , and \mathcal{M}

- (i) We define \mathcal{R} as the set of all restricted (d, k) sequences containing only **one** occurrence of w , located at the **right end**.
- (ii) We also define \mathcal{U} as

$$\mathcal{U} = \{u : w \cdot u \in \mathcal{T}_1\},$$

that is, a word $u \in \mathcal{U}$ if u is a restricted (d, k) sequence and $w \cdot u$ has exactly **one** occurrence of w at the **left end** of $w \cdot u$.

- (iii) \mathcal{M} is defined as

$$\mathcal{M} = \{u : w \cdot u \in \mathcal{T}_2 \text{ and } w \text{ occurs at the right of } w \cdot u\},$$

that is, \mathcal{M} is a language such that any word in $\{w\} \cdot \mathcal{M}$ has exactly **two** occurrences of w at the **left and right ends**.

Example: Let $w = 0100101$. Notice $010100101 \in \mathcal{R}$, and $01 \in \mathcal{U}$.
Observe $00101 \notin \mathcal{U}$, but $00101 \in \mathcal{M}$ because $010010100101 \in \mathcal{T}_2$.

Language Relationships and Generating Functions

The following holds:

$$\begin{aligned}
 \mathcal{T}_r &= \mathcal{R} \cdot \mathcal{M}^{r-1} \cdot \mathcal{U} & \mathcal{M}^* &= \mathcal{B}_{d,k}^* \cdot \{w\} + \mathcal{S} \\
 \mathcal{T}_0 \cdot \{w\} &= \mathcal{R} \cdot \mathcal{S} & \mathcal{U} \cdot \mathcal{B}_{d,k} &= \mathcal{M} + \mathcal{U} - \{\epsilon\} \\
 & & \{w\} \cdot \mathcal{M} &= \mathcal{B}_{d,k} \cdot \mathcal{R} - (\mathcal{R} - \{w\})
 \end{aligned}$$

Then, the above language relationships translate into

$$\frac{1}{1 - M(z)} = \frac{1}{1 - B(z)} \cdot z^m P(w) + S(z),$$

$$U(z) = \frac{M(z) - 1}{B(z) - 1}, \quad R(z) = z^m P(w) \cdot U(z)$$

where $P(w)$ is the probability of w , and m is the length of w .

In particular, we find

$$T_0(z) = \frac{S(z)}{D(z)}, \quad T_r(z) = \frac{z^m P(w) (D(z) + B(z) - 1)^{r-1}}{D(z)^{r+1}},$$

where $S(z)$ is the autocorrelation polynomial for w and

$$D(z) = S(z)(1 - B(z)) + z^m P(w).$$

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Number of Occurrences

Let O_n be a random variable representing the number of occurrences of w in a (regular) binary sequence of length n .

The probability generating function of T_r ,

$$T_r(z) = \sum_{n \geq 0} P(O_n = r, \mathcal{D}_n) z^n,$$

where

$\mathcal{D}_n =$ the event that a randomly generated binary sequence of length n is a (d, k) sequence.

Define the bivariate generating function as

$$T(z, u) = \sum_{r \geq 0} T_r(z) u^r = \sum_{r \geq 0} \sum_{n \geq 0} P(O_n = r, \mathcal{D}_n) z^n u^r.$$

The probability that a randomly generated sequence of length n is a (d, k) sequence is

$$P(\mathcal{D}_n) = [z^n] T(z, 1).$$

Number of Occurrences

Introduce a short-hand notation $O_n(\mathcal{D}_n)$ for the conditional number of occurrences of w in a (d, k) sequence,

$$P(O_n(\mathcal{D}_n) = r) = P(O_n = r \mid \mathcal{D}_n).$$

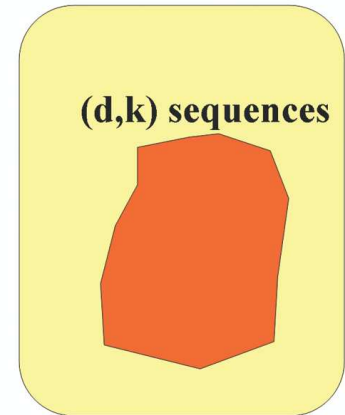
The probability generating function of $O_n(\mathcal{D}_n)$,

$$\mathbf{E}[u^{O_n(\mathcal{D}_n)}] = \frac{[z^n]T(z, u)}{[z^n]T(z, 1)}.$$

The mean and second factorial moment of $O_n(\mathcal{D}_n)$ can be computed by

$$\mathbf{E}[O_n(\mathcal{D}_n)] = \frac{[z^n]T_u(z, 1)}{[z^n]T(z, 1)}, \quad \mathbf{E}[O_n(\mathcal{D}_n)(O_n(\mathcal{D}_n) - 1)] = \frac{[z^n]T_{uu}(z, 1)}{[z^n]T(z, 1)}.$$

Binary sequences



Main Results

Theorem 1. Let $\rho := \rho(p) = 1/\lambda$ be the unique positive real root of

$$1 - B(z) = 0.$$

Then

$$P(\mathcal{D}_n) = \frac{1}{B'(\rho)} \lambda^{n+1} + O(\omega^n)$$

is the *probability of generating a (d, k) sequence* for some $\omega < \lambda$.
Furthermore, the *mean* is

$$\mathbf{E}[O_n(\mathcal{D}_n)] = \frac{(n - m + 1)P(w)}{B'(\rho)} \lambda^{-m+1} + O(1),$$

and the *variance* becomes

$$\begin{aligned} \mathbf{Var}[O_n(\mathcal{D}_n)] &= (n - m + 1)P(w) \left[\frac{(1 - 2m)P(w)}{B'(\rho)^2} \lambda^{-2m+2} \right. \\ &\quad \left. + \frac{P(w)B''(\rho)}{B'(\rho)^3} \lambda^{-2m+1} + \frac{2S(\rho) - 1}{B'(\rho)} \lambda^{-m+1} \right] + O(1). \end{aligned}$$

Main Results

Theorem 2. Let $\tau := \tau(p, w)$ be the smallest real root of

$$D(z) = 0, \quad (\text{cf. } D(z) = S(z)(1 - B(z)) + z^m P(w))$$

and $\rho := \rho(p)$ be the unique positive real root of $B(z) = 1$.

(i) For $r = O(1)$,

$$P(O_n(\mathcal{D}_n) = r) \sim \frac{P(w)B'(\rho)(1 - B(\tau))^{r-1}}{D'(\tau)^{r+1}\tau^{r-m}} \binom{n - m + r}{r} \left(\frac{\rho}{\tau}\right)^{n+1}$$

for large n and $r \geq 1$.

(ii) (Central limit) For $r = \mathbf{E}[O_n(\mathcal{D}_n)] + x\sqrt{\mathbf{Var}[O_n(\mathcal{D}_n)]}$ with $x = O(1)$,

$$\frac{O_n(\mathcal{D}_n) - \mathbf{E}[O_n(\mathcal{D}_n)]}{\sqrt{\mathbf{Var}[O_n(\mathcal{D}_n)]}} \xrightarrow{d} N(0, 1)$$

where $N(0, 1)$ is the standard normal distribution.

Main Results

(iii) (Large deviations) For $r = (1 + \delta)\mathbf{E}[O_n(\mathcal{D}_n)]$ with $\delta > 0$, let a be a real constant such that

$$na = (1 + \delta)\mathbf{E}[O_n(\mathcal{D}_n)]$$

and let

$$h_a(z) = a \log M(z) - \log z.$$

Let also z_a be the unique real root of the equation $h'_a(z) = 0$ such that $z_a \in (0, \rho)$. Then,

$$P(O_n(D_n) = na) = \frac{c_1 \cdot e^{-nI(a)}}{\sqrt{2\pi n}} \left(1 + \frac{c_2}{n} + O\left(\frac{1}{n^2}\right) \right)$$

and

$$P(O_n(D_n) \geq na) = \frac{c_1 \cdot e^{-nI(a)}}{\sqrt{2\pi n}(1 - M(z_a))} \left(1 + O\left(\frac{1}{n}\right) \right)$$

where

$$I(a) = -\log \rho - h_a(z_a),$$

and the constants c_1 and c_2 are explicitly computable.

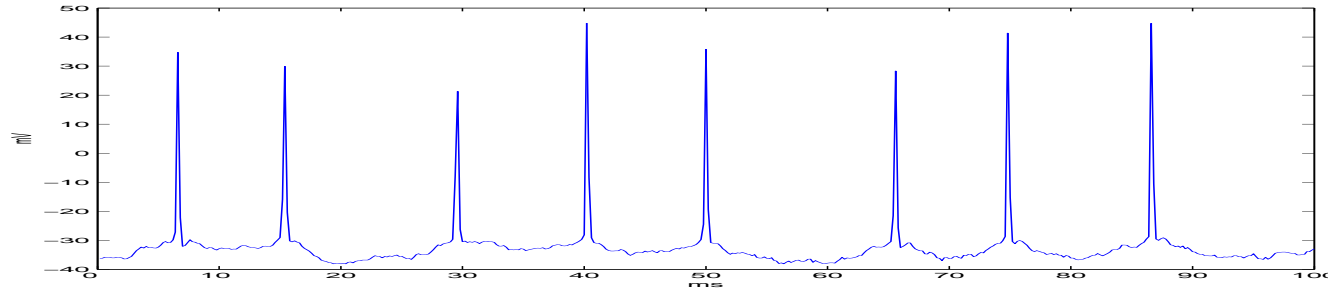
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Experimental Results

Spike trains of neuronal data satisfy structural constraints that exactly match the framework of (d, k) binary sequences.

spike train :



(d, k) sequence : 010001000000100001000100000010001000001000010000010000...

Question: How can we classify a pattern as significant?

We use the large deviations results to detect under- and over-represented patterns.

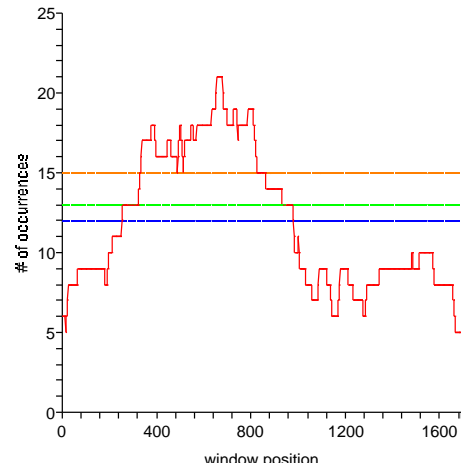
The threshold, O_{th} , above which pattern occurrences will be classified as statistically significant, is defined as the minimum O_{th} such that

$$P(O_n(\mathcal{D}_n) \geq O_{th}) \leq \alpha_{th}$$

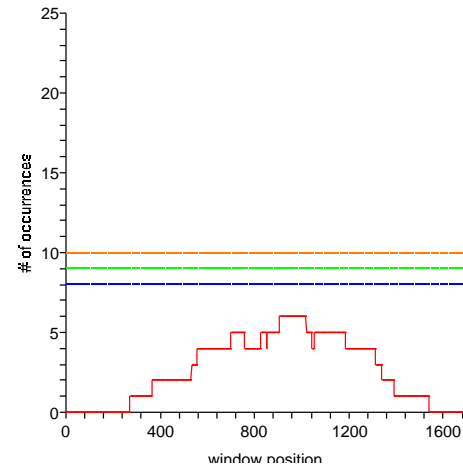
where α_{th} is a given probability threshold (e.g. $\alpha_{th} = 10^{-6}, 10^{-8}$).

Experimental Results

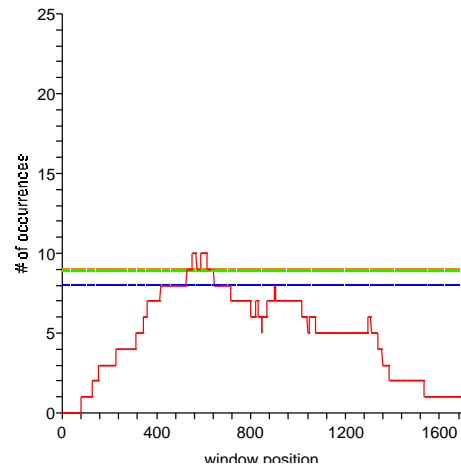
Number of occurrences of w within a window of size 500; here $[i] = \underbrace{0 \dots 0}_{i-1} 1$.



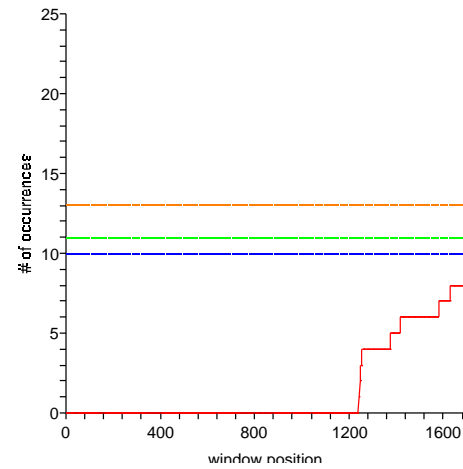
(a) $w=(4)(4)(4)$



(b) $w=(5)(3)(5)$



(c) $w=(4)(5)(3)$



(d) $w=(5)(5)(5)$

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Analysis : Large Deviation Result

Theorem For $r = (1 + \delta)\mathbf{E}[O_n(\mathcal{D}_n)]$ with $\delta > 0$, let a be a real constant such that

$$na = (1 + \delta)\mathbf{E}[O_n(\mathcal{D}_n)]$$

and let

$$h_a(z) = a \log M(z) - \log z.$$

Let also z_a be the unique real root of the equation $h'_a(z) = 0$ such that $z_a \in (0, \rho)$. Then,

$$P(O_n(D_n) = na) = \frac{c_1 \cdot e^{-nI(a)}}{\sqrt{2\pi n}} \left(1 + \frac{c_2}{n} + O\left(\frac{1}{n^2}\right) \right)$$

where

$$I(a) = -\log \rho - h_a(z_a),$$

and the constants c_1 and c_2 are explicitly computable.

Analysis : Sketch of the Proof

1. Generating functions and Cauchy coefficient formula

$$P(O_n(D_n) = na) = [u^{na}]T_n(u) = \frac{[z^n][u^{na}]T(z, u)}{[z^n]T(z, 1)} = \frac{[z^n][u^{na}]T(z, u)}{P(\mathcal{D}_n)}$$

$$[u^{na}]T(z, u) = \frac{P(w)z^m}{D(z)^2} M(z)^{na-1}$$

$$\begin{aligned} [z^n][u^{na}]T(z, u) &= \frac{1}{2\pi i} \oint \frac{P(w)z^m}{D(z)^2} M(z)^{na-1} \frac{1}{z^{n+1}} dz \\ &= \frac{1}{2\pi i} \oint e^{nh_a(z)} g(z) dz \end{aligned}$$

where

$$h_a(z) = a \log M(z) - \log z \quad \text{and} \quad g(z) = \frac{P(w)z^{m-1}}{D(z)^2 M(z)}.$$

Analysis : Sketch of the Proof

2. Saddle point contour

Let z_a a unique real root of the equation $h'_a(z) = 0$. We evaluate the integral on $\mathcal{C} = \{z : |z| = z_a\}$

3. Contour split

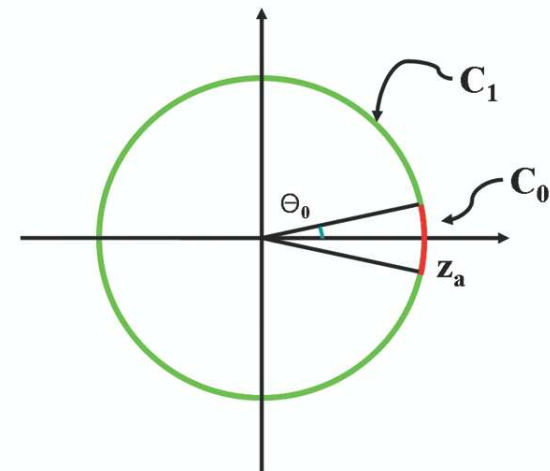
We split \mathcal{C} into \mathcal{C}_0 and \mathcal{C}_1 where

$$\mathcal{C}_0 = \{z \in \mathcal{C} : |\arg(z)| \leq \theta_0\}$$

and

$$\mathcal{C}_1 = \{z \in \mathcal{C} : |\arg(z)| \geq \theta_0\}$$

for $\theta_0 = n^{-2/5}$.



$$\begin{aligned} & [z^n][u^{na}]T(z, u) \\ &= I_0 + I_1 \\ &= \frac{1}{2\pi i} \int_{\mathcal{C}_0} e^{nh_a(z)} g(z) dz + \frac{1}{2\pi i} \int_{\mathcal{C}_1} e^{nh_a(z)} g(z) dz. \end{aligned}$$

Analysis : Sketch of the Proof

4. Approximation of I_0

Using **change of variables** and **Taylor series expansion**, we get

$$\begin{aligned} I_0 &= \frac{1}{2\pi i} \int_{c_0} e^{nh_a(z)} g(z) dz = \frac{1}{2\pi} \int_{-\theta_0}^{+\theta_0} e^{nh_a(z_a e^{i\theta})} g(z_a e^{i\theta}) z_a e^{i\theta} d\theta \\ &\sim \frac{e^{nh_a(z_a)}}{2\pi \tau_a \sqrt{n}} \int_{-\infty}^{+\infty} \exp\left(-\frac{\omega^2}{2}\right) F(\omega) d\omega = \frac{g(z_a) e^{nh_a(z_a)}}{\tau_a \sqrt{2\pi n}} \left(1 + \frac{c_2}{n} + O\left(\frac{1}{n^2}\right)\right) \end{aligned}$$

5. Elimination of I_1

We show that I_1 is **exponentially smaller** than I_0 .

$M(z)$ is the probability generating function of language \mathcal{M} . By its **non-negativity of coefficients** and **aperiodicity**, $|M(z_a e^{i\theta})|$ is **uniquely maximum** at $\theta = 0$. For $\theta \in [\theta_0, \pi]$,

$$\left| e^{nh_a(z_a e^{i\theta})} \right| = \frac{|M(z_a e^{i\theta})|^{na}}{z_a^n} \leq \frac{|M(z_a e^{i\theta_0})|^{na}}{z_a^n} = \left| e^{nh_a(z_a e^{i\theta_0})} \right|.$$

Analysis : Sketch of the Proof

6. Putting together

$$\begin{aligned} P(O_n(D_n) = na) &= \frac{[z^n][u^{na}]T(z, u)}{[z^n]T(z, 1)} = \frac{I_0 + I_1}{P(\mathcal{D}_n)} = \frac{I_0 \left(1 + O\left(e^{-cn^{1/5}}\right)\right)}{P(\mathcal{D}_n)} \\ &= \frac{c_1 \cdot e^{-nI(a)}}{\sqrt{2\pi n}} \left(1 + \frac{c_2}{n} + O\left(\frac{1}{n^2}\right)\right) \end{aligned}$$

where

$$I(a) = -\log \rho - h_a(z_a).$$